Package 'PAGFL'

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Title Joint Estimation of Latent Groups and Group-Specific Coefficients in Panel Data Models

Version 1.1.2

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Description Latent group structures are a common challenge in panel data analysis. Disregarding group-level heterogeneity can introduce bias. Conversely, estimating individual coefficients for each cross-sectional unit is inefficient and may lead to high uncertainty. This package addresses the issue of unobservable group structures by implementing the pairwise adaptive group fused Lasso (PAGFL) by Mehrabani (2023) [<doi:10.1016/j.jeconom.2022.12.002>](https://doi.org/10.1016/j.jeconom.2022.12.002). PAGFL identifies latent group structures and group-specific coefficients in a single step. On top of that, we extend the PAGFL to time-varying coefficient functions.

License $AGPL (> = 3)$

Encoding UTF-8

RoxygenNote 7.3.2

LinkingTo Rcpp, RcppArmadillo, RcppParallel

Imports Rcpp, lifecycle, ggplot2, RcppParallel

BugReports <https://github.com/Paul-Haimerl/PAGFL/issues>

URL <https://github.com/Paul-Haimerl/PAGFL>

Suggests test that $(>= 3.0.0)$

Config/testthat/edition 3

NeedsCompilation yes

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Depends $R (= 3.5.0)$

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Contents

grouped_plm *Grouped Panel Data Model*

Description

Estimate a grouped panel data model given an observed group structure. Slope parameters are homogeneous within groups but heterogeneous across groups. This function supports both static and dynamic panel data models, with or without endogenous regressors.

```
grouped_plm(
  formula,
  data,
  groups,
  index = NULL,
  n_periods = NULL,
  method = "PLS",
  Z = NULL,bias_correc = FALSE,
  rho = 0.07 \times \log(N \times n_{\text{periods}})/sqrt(N \times n_{\text{periods}}),
  verbose = TRUE,
  parallel = TRUE,
  ...
\mathcal{E}## S3 method for class 'gplm'
print(x, \ldots)## S3 method for class 'gplm'
formula(x, \ldots)## S3 method for class 'gplm'
df.residual(object, ...)
## S3 method for class 'gplm'
summary(object, ...)
```

```
## S3 method for class 'gplm'
coef(object, ...)
## S3 method for class 'gplm'
residuals(object, ...)
## S3 method for class 'gplm'
fitted(object, ...)
```


Details

Consider the grouped panel data model

$$
y_{it} = \gamma_i + \beta'_i x_{it} + \epsilon_{it}, \quad i = 1, \dots, N, t = 1, \dots, T,
$$

where y_{it} is the scalar dependent variable, γ_i is an individual fixed effect, x_{it} is a $p \times 1$ vector of explanatory variables, and ϵ_{it} is a zero mean error. The coefficient vector β_i is subject to the observed group pattern

$$
\beta_i = \sum_{k=1}^K \alpha_k \mathbf{1} \{ i \in G_k \},
$$

with $\bigcup_{k=1}^{K} G_k = \{1, \ldots, N\}, G_k \cap G_j = \emptyset$ and $\|\alpha_k - \alpha_j\| \neq 0$ for any $k \neq j, k = 1, \ldots, K$.

Using *PLS*, the group-specific coefficients for group k are obtained via *OLS*

$$
\hat{\alpha}_k = \left(\sum_{i \in G_k} \sum_{t=1}^T \tilde{x}_{it} \tilde{x}'_{it}\right)^{-1} \sum_{i \in G_k} \sum_{t=1}^T \tilde{x}_{it} \tilde{y}_{it},
$$

where $\tilde{a}_{it} = a_{it} - T^{-1} \sum_{t=1}^{T} a_{it}$, $a = \{y, x\}$ to concentrate out the individual fixed effects γ_i (within-transformation).

In case of *PGMM*, the slope coefficients are derived as

$$
\hat{\alpha}_k = \left(\left[\sum_{i \in G_k} T^{-1} \sum_{t=1}^T z_{it} \Delta x_{it} \right]' W_k \left[\sum_{i \in G_k} T^{-1} \sum_{t=1}^T z_{it} \Delta x_{it} \right] \right)^{-1}
$$

$$
\left[\sum_{i \in G_k} T^{-1} \sum_{t=1}^T z_{it} \Delta x_{it} \right]' W_k \left[\sum_{i \in G_k} T^{-1} \sum_{t=1}^T z_{it} \Delta y_{it} \right],
$$

where W_k is a $q \times q$ p.d. symmetric weight matrix and Δ denotes the first difference operator $\Delta x_{it} = x_{it} - x_{it-1}$ (first-difference transformation).

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Value

An object of class gplm holding

A gplm object has print, summary, fitted, residuals, formula, df.residual, and coef S3 methods.

Author(s)

Paul Haimerl

References

Dhaene, G., & Jochmans, K. (2015). Split-panel jackknife estimation of fixed-effect models. *The Review of Economic Studies*, 82(3), 991-1030. [doi:10.1093/restud/rdv007.](https://doi.org/10.1093/restud/rdv007) Mehrabani, A. (2023). Estimation and identification of latent group structures in panel data. *Journal of Econometrics*, 235(2), 1464-1482. [doi:10.1016/j.jeconom.2022.12.002.](https://doi.org/10.1016/j.jeconom.2022.12.002)

Examples

```
# Simulate a panel with a group structure
sim < sum\_DGP(N = 20, n\_periods = 80, p = 2, n\_groups = 3)y <- sim$y
X \leftarrow \text{sim$} \groups <- sim$groups
df \leftarrow \text{cbind}(y = c(y), X)# Estimate the grouped panel data model
estim \le grouped_plm(y \sim ., data = df, groups = groups, n_periods = 80, method = "PLS")
summary(estim)
# Lets pass a panel data set with explicit cross-sectional and time indicators
i_index <- rep(1:20, each = 80)
t_{index} < - rep(1:80, 20)df \le data.frame(y = c(y), X, i_index = i_index, t_index = t_index)
estim <- grouped_plm(
  y \sim .data = df, index = c("i_index", "t_index"), groups = groups, method = "PLS"\lambdasummary(estim)
```


Description

Estimate a grouped time-varying panel data model given an observed group structure. Coefficient functions are homogeneous within groups but heterogeneous across groups. The time-varying coefficients are modeled as polynomial B-splines. The function supports both static and dynamic panel data models.

```
grouped_tv_plm(
  formula,
  data,
  groups,
  index = NULL,n_periods = NULL,
 d = 3,
 M = \text{floor}(\text{length}(y) \cdot (1/7) - \text{log}(p)),const_coef = NULL,
  rho = 0.04 * log(N * n\_periods)/sqrt(N * n\_periods),
  verbose = TRUE,
 parallel = TRUE,
  ...
)
## S3 method for class 'tv_gplm'
summary(object, ...)
## S3 method for class 'tv_gplm'
formula(x, \ldots)## S3 method for class 'tv_gplm'
df.residual(object, ...)
## S3 method for class 'tv_gplm'
print(x, \ldots)## S3 method for class 'tv_gplm'
coef(object, ...)
## S3 method for class 'tv_gplm'
residuals(object, ...)
## S3 method for class 'tv_gplm'
fitted(object, ...)
```
Arguments

Details

Consider the grouped time-varying panel data model

$$
y_{it} = \gamma_i + \beta_i'(t/T)x_{it} + \epsilon_{it}, \quad i = 1, \dots, N, t = 1, \dots, T,
$$

where y_{it} is the scalar dependent variable, γ_i is an individual fixed effect, x_{it} is a $p \times 1$ vector of explanatory variables, and ϵ_{it} is a zero mean error. The coefficient vector $\beta_i(t/T)$ is subject to the observed group pattern \overline{K}

$$
\beta_i\left(\frac{t}{T}\right) = \sum_{k=1}^R \alpha_k\left(\frac{t}{T}\right) \mathbf{1}\{i \in G_k\},\
$$

with $\bigcup_{k=1}^{K} G_k = \{1, \ldots, N\}, G_k \cap G_j = \emptyset$ and $\|\alpha_k - \alpha_j\| \neq 0$ for any $k \neq j, k = 1, \ldots, K$.

 $\alpha_k(t/T)$ and, in turn, $\beta_i(t/T)$ is estimated as polynomial B-splines using the penalized sievetechnique. To this end, let $B(v)$ denote a $M + d + 1$ vector of polynomial spline basis functions, where d represents the polynomial degree and M gives the number of interior knots of the Bspline. $\alpha_k(t/T)$ is approximated by forming a linear combination of the basis functions $\alpha_k(t/T) \approx$ $\xi_k' B(t/T)$, where ξ_k is a $(M + d + 1) \times p$ coefficient matrix.

The explanatory variables are projected onto the spline basis system, which results in the $(M + d +$ $1)p \times 1$ vector $z_{it} = x_{it} \otimes B(v)$. Subsequently, the DGP can be reformulated as

$$
y_{it} = \gamma_i + z'_{it} \text{vec}(\pi_i) + u_{it},
$$

where $\pi_i = \xi_k$ if $i \in G_k$, $u_{it} = \epsilon_{it} + \eta_{it}$, and η_{it} reflects a sieve approximation error. We refer to Su et al. (2019, sec. 2) for more details on the sieve technique.

Finally, $\hat{\alpha}_k(t/T)$ is obtained as $\hat{\alpha}_k(t/T) = \hat{\xi}'_k B(t/T)$, where the vector of control points ξ_k is estimated using *OLS*

$$
\hat{\xi}_k = \left(\sum_{i \in G_k} \sum_{t=1}^T \tilde{z}_{it} \tilde{z}'_{it}\right)^{-1} \sum_{i \in G_k} \sum_{t=1}^T \tilde{z}_{it} \tilde{y}_{it},
$$

and $\tilde{a}_{it} = a_{it} - T^{-1} \sum_{t=1}^{T} a_{it}$, $a = \{y, z\}$ to concentrate out the fixed effect γ_i (within-transformation). In case of an unbalanced panel data set, the earliest and latest available observations per group define the start and end-points of the interval on which the group-specific time-varying coefficients are defined.

Value

An object of class tv_gplm holding

An object of class tv_gplm has print, summary, fitted, residuals, formula, df.residual and coef S3 methods.

Author(s)

Paul Haimerl

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References

Su, L., Wang, X., & Jin, S. (2019). Sieve estimation of time-varying panel data models with latent structures. *Journal of Business & Economic Statistics*, 37(2), 334-349. [doi:10.1080/07350015.2017.1340299.](https://doi.org/10.1080/07350015.2017.1340299)

Examples

```
# Simulate a time-varying panel with a trend and a group pattern
set.seed(1)
sim < -sim_{av}tv_DGP(N = 10, n_periods = 50, intercept = TRUE, p = 2)
df \leftarrow data-frame(y = c(sim\{y}))
groups <- sim$groups
# Estimate the time-varying grouped panel data model
estim \le grouped_tv_plm(y \sim ., data = df, n_periods = 50, groups = groups)
summary(estim)
```
pagfl *Pairwise Adaptive Group Fused Lasso*

Description

Estimate panel data models with a latent group structure using the pairwise adaptive group fused Lasso (*PAGFL*) by Mehrabani (2023). The *PAGFL* jointly identifies the group structure and groupspecific slope parameters. The function supports both static and dynamic panels, with or without endogenous regressors.

```
pagfl(
  formula,
  data,
  index = NULL,n_periods = NULL,
  lambda,
 method = "PLS",Z = NULL,min_group_frac = 0.05,
 bias_correc = FALSE,
  kappa = 2,
 max\_iter = 5000,
  tol_convergence = 1e-08,
  tol\_group = 0.001,rho = 0.07 \times \log(N \times n periods)/sqrt(N \times n periods),
  varrho = max(sqrt(5 * N * n_periods * p)/log(N * n_periods * p) - 7, 1),
  verbose = TRUE,
  parallel = TRUE,
```

```
...
\mathcal{L}## S3 method for class 'pagfl'
print(x, \ldots)## S3 method for class 'pagfl'
formula(x, ...)
## S3 method for class 'pagfl'
df.residual(object, ...)
## S3 method for class 'pagfl'
summary(object, ...)
## S3 method for class 'pagfl'
coef(object, ...)
## S3 method for class 'pagfl'
residuals(object, ...)
## S3 method for class 'pagfl'
fitted(object, ...)
```


"PGMM" for using the penalized Generalized Method of Moments (*PGMM*). *PGMM* is required when instrumenting endogenous regressors, in which case a matrix Z containing the necessary exogenous instruments must be supplied (Mehrabani, 2023, sec. 2.3).

Default is "PLS".

- Z a $NT \times q$ matrix or data.frame of exogenous instruments, where $q \geq p$, $\mathbf{Z} = (z_1, \ldots, z_N)'$, $z_i = (z_{i1}, \ldots, z_{iT})'$ and z_{it} is a $q \times 1$ vector. Z is only required when method = "PGMM" is selected. When using "PLS", either pass NULL or Z is disregarded. Default is NULL.
- min_group_frac the minimum group cardinality as a fraction of the total number of individuals N. In case a group falls short of this threshold, each of its members is allocated to one of the remaining groups according to the *MSE*. Default is 0.05.
- bias_correc logical. If TRUE, a Split-panel Jackknife bias correction following Dhaene and Jochmans (2015) is applied to the slope parameters. We recommend using the correction when working with dynamic panels. Default is FALSE.
- kappa the a non-negative weight used to obtain the adaptive penalty weights. Default is 2.
- max_{iter} the maximum number of iterations for the *ADMM* estimation algorithm. Default is $1 * 10^4$.

tol_convergence

- the tolerance limit for the stopping criterion of the iterative *ADMM* estimation algorithm. Default is $1 * 10^{-8}$.
- tol_group the tolerance limit for within-group differences. Two individuals i, j are assigned to the same group if the Frobenius norm of their coefficient vector difference is below this threshold. Default is $1 * 10^{-3}$.
- rho the tuning parameter balancing the fitness and penalty terms in the IC that determines the penalty parameter λ . If left unspecified, the heuristic $\rho = 0.07 \frac{\log (NT)}{\sqrt{NT}}$ of Mehrabani (2023, sec. 6) is used. We recommend the default.
- varrho the non-negative Lagrangian $ADMM$ penalty parameter. For *PLS*, the ρ value is trivial. However, for *PGMM*, small values lead to slow convergence. If left unspecified, the default heuristic $\rho = \max(\frac{\sqrt{5NTp}}{\log(NTp)} - 7, 1)$ is used.
- verbose logical. If TRUE, helpful warning messages are shown. Default is TRUE.

parallel logical. If TRUE, certain operations are parallelized across multiple cores. Default is TRUE. ellipsis x of class pagfl.

object of class pagfl.

Details

Consider the grouped panel data model

$$
y_{it} = \gamma_i + \beta'_i x_{it} + \epsilon_{it}, \quad i = 1, \dots, N, t = 1, \dots, T,
$$

where y_{it} is the scalar dependent variable, γ_i is an individual fixed effect, x_{it} is a $p \times 1$ vector of weakly exogenous explanatory variables, and ϵ_{it} is a zero mean error. The coefficient vector β_i is subject to the latent group pattern

$$
\beta_i = \sum_{k=1}^K \alpha_k \mathbf{1} \{ i \in G_k \},
$$

with $\bigcup_{k=1}^{K} G_k = \{1, \ldots, N\}, G_k \cap G_j = \emptyset$ and $\|\alpha_k - \alpha_j\| \neq 0$ for any $k \neq j, k = 1, \ldots, K$.

The PLS method jointly estimates the latent group structure and group-specific coefficients by minimizing the criterion

$$
Q_{NT}(\beta, \lambda) = \frac{1}{T} \sum_{i=1}^{N} \sum_{t=1}^{T} (\tilde{y}_{it} - \beta_i' \tilde{x}_{it})^2 + \frac{\lambda}{N} \sum_{i=1}^{N-1} \sum_{j>i}^{N} \dot{\omega}_{ij} ||\beta_i - \beta_j||
$$

with respect to $\boldsymbol{\beta} = (\beta'_1, \dots, \beta'_N)'$. $\tilde{a}_{it} = a_{it} - T^{-1} \sum_{t=1}^T a_{it}$, $a = \{y, x\}$ to concentrate out the individual fixed effects γ_i . λ is the penalty tuning parameter and $\dot{\omega}_{ij}$ reflects adaptive penalty weights (see Mehrabani, 2023, eq. 2.6). $\|\cdot\|$ denotes the Frobenius norm. The adaptive weights \dot{w}_{ij} are obtained by a preliminary individual least squares estimation. The criterion function is minimized via an iterative alternating direction method of multipliers (*ADMM*) algorithm (see Mehrabani, 2023, sec. 5.1).

PGMM employs a set of instruments Z to control for endogenous regressors. Using *PGMM*, β is estimated by minimizing

$$
Q_{NT}(\boldsymbol{\beta}, \lambda) = \sum_{i=1}^{N} \left[\frac{1}{N} \sum_{t=1}^{T} z_{it} (\Delta y_{it} - \beta_i' \Delta x_{it}) \right]' W_i \left[\frac{1}{T} \sum_{t=1}^{T} z_{it} (\Delta y_{it} - \beta_i' \Delta x_{it}) \right] + \frac{\lambda}{N} \sum_{i=1}^{N-1} \sum_{j>i}^{N} \ddot{\omega}_{ij} ||\beta_i - \beta_j||.
$$

 $\ddot{\omega}_{ij}$ are obtained by an initial *GMM* estimation. Δ gives the first differences operator Δy_{it} = $y_{it} - y_{it-1}$. W_i represents a data-driven $q \times q$ weight matrix. I refer to Mehrabani (2023, eq. 2.10) for more details. Again, the criterion function is minimized using an efficient *ADMM* algorithm (Mehrabani, 2023, sec. 5.2).

Two individuals are assigned to the same group if $\|\hat{\beta}_i - \hat{\beta}_j\| \leq \epsilon_{\text{tol}}$, where ϵ_{tol} is determined by tol_group. Subsequently, the number of groups follows as the number of distinct elements in β . Given an estimated group structure, it is straightforward to obtain post-Lasso estimates using group-wise least squares or *GMM* (see [grouped_plm](#page-1-1)).

We recommend identifying a suitable λ parameter by passing a logarithmically spaced grid of candidate values with a lower limit close to 0 and an upper limit that leads to a fully homogeneous panel. A BIC-type information criterion then selects the best fitting λ value.

Value

An object of class pagfl holding

model a data.frame containing the dependent and explanatory variables as well as cross-sectional and time indices,

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A pagfl object has print, summary, fitted, residuals, formula, df.residual, and coef S3 methods.

Author(s)

Paul Haimerl

References

Dhaene, G., & Jochmans, K. (2015). Split-panel jackknife estimation of fixed-effect models. *The Review of Economic Studies*, 82(3), 991-1030. [doi:10.1093/restud/rdv007.](https://doi.org/10.1093/restud/rdv007) Mehrabani, A. (2023). Estimation and identification of latent group structures in panel data. *Journal of Econometrics*, 235(2), 1464-1482. [doi:10.1016/j.jeconom.2022.12.002.](https://doi.org/10.1016/j.jeconom.2022.12.002)

Examples

```
# Simulate a panel with a group structure
sim < -sim\_DGP(N = 20, n\_periods = 80, p = 2, n\_groups = 3)y \le -\sin\frac{\pi}{3}yX \leftarrow \text{sim$} \df \leftarrow \text{cbind}(y = c(y), X)# Run the PAGFL procedure
estim \le pagfl(y \sim ., data = df, n_periods = 80, lambda = 0.5, method = "PLS")
summary(estim)
# Lets pass a panel data set with explicit cross-sectional and time indicators
i<sub>n</sub>index <- rep(1:20, each = 80)
t_{index} < - rep(1:80, 20)df \leq - data.frame(y = c(y), X, i_index = i_index, t_index = t_index)
estim <- pagfl(
  y \sim ., data = df, index = c("i_index", "t_index"), lambda = 0.5, method = "PLS"
)
summary(estim)
```


Description

Construct a static or dynamic, exogenous or endogenous panel data set subject to a group structure in the slope coefficients with optional $AR(1)$ or $GARCH(1, 1)$ innovations.

Usage

```
sim_DGP(
 N = 50,
 n_periods = 40,
 p = 2,n_{\text{g}roups} = 3,
  group_proportions = NULL,
 error_spec = "iid",
  dynamic = FALSE,
  dyn_panel = lifecycle::deprecated(),
 q = NULL,alpha_0 = NULL\mathcal{L}
```


Details

The scalar dependent variable y_{it} is generated according to the following grouped panel data model

$$
y_{it} = \gamma_i + \beta'_i x_{it} + u_{it}, \quad i = \{1, ..., N\}, \quad t = \{1, ..., T\}.
$$

 γ_i represents individual fixed effects and x_{it} a $p \times 1$ vector of regressors. The individual slope coefficient vectors β_i are subject to a group structure

$$
\beta_i = \sum_{k=1}^K \alpha_k \mathbf{1} \{ i \in G_k \},
$$

with $\cup_{k=1}^K G_k = \{1, \ldots, N\}$, $G_k \cap G_j = \emptyset$ and $\|\alpha_k - \alpha_j\| \neq 0$ for any $k \neq j$, $k = 1, \ldots, K$. The total number of groups K is determined by n_groups.

If a panel data set with exogenous regressors is generated (set $q = NULL$), the explanatory variables are simulated according to

$$
x_{it,j} = 0.2\gamma_i + e_{it,j}, \quad \gamma_i, e_{it,j} \sim i.i.d.N(0,1), \quad j = \{1,\ldots,p\},\
$$

where $e_{it,j}$ denotes a series of innovations. γ_i and e_i are independent of each other.

In case alpha_0 = NULL, the group-level slope parameters α_k are drawn from $\sim U[-2, 2]$.

If a dynamic panel is specified (dynamic = TRUE), the AR coefficients β_i^{AR} are drawn from a uniform distribution with support $(-1, 1)$ and $x_{it,j} = e_{it,j}$. Moreover, the individual fixed effects enter the dependent variable via $(1 - \beta_i^{AR})\gamma_i$ to account for the autoregressive dependency. We refer to Mehrabani (2023, sec 6) for details.

When specifying an endogenous panel (set q to $q \ge p$), the $e_{it,j}$ correlate with the cross-sectional innovations u_{it} by a magnitude of 0.5 to produce endogenous regressors ($E(u|X) \neq 0$). However, the endogenous regressors can be accounted for by exploiting the q instruments in Z , for which $E(u|Z) = 0$ holds. The instruments and the first stage coefficients are generated in the same fashion as X and α when q = NULL.

The function nests, among other, the DGPs employed in the simulation study of Mehrabani (2023, sec. 6).

Value

A list holding

Author(s)

Paul Haimerl

References

Mehrabani, A. (2023). Estimation and identification of latent group structures in panel data. *Journal of Econometrics*, 235(2), 1464-1482. [doi:10.1016/j.jeconom.2022.12.002.](https://doi.org/10.1016/j.jeconom.2022.12.002)

Examples

```
# Simulate DGP 1 from Mehrabani (2023, sec. 6)
alpha_0DGP1 <- matrix(c(0.4, 1, 1.6, 1.6, 1, 0.4), ncol = 2)
DGP1 <- sim_DGP(
  N = 50, n_periods = 20, p = 2, n_groups = 3,
  group_proportions = c(.4, .3, .3), alpha_0 = alpha_0_DGP1
\mathcal{L}
```


Simulate a Time-varying Panel With a Group Structure in the Slope *Coefficients*

Description

Construct a time-varying panel data set subject to a group structure in the slope coefficients with optional $AR(1)$ innovations.

```
sim_tv_DGP(
 N = 50,
  n_periods = 40,
  intercept = TRUE,
 p = 1,
  n_{\text{g} \cdot n} = 3,
  d = 3,
  dynamic = FALSE,
  group_proportions = NULL,
```
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```
error\_spec = "iid",locations = NULL,
 scales = NULL,
 polynomial_coef = NULL,
 sd\_error = 1,
 DGP = lifecycle::deprecated()
\mathcal{L}
```


Details

The scalar dependent variable y_{it} is generated according to the following time-varying grouped panel data model

$$
y_{it} = \gamma_i + \beta'_{it} x_{it} + u_{it}, \quad i = 1, ..., N, t = 1, ..., T,
$$

where γ_i is an individual fixed effect and x_{it} is a $p \times 1$ vector of explanatory variables. The coefficient vector $\beta_i = {\beta'_{i1}, \dots, \beta'_{iT}}'$ is subject to the group pattern

$$
\beta_i\left(\frac{t}{T}\right) = \sum_{k=1}^K \alpha_k\left(\frac{t}{T}\right) \mathbf{1}\{i \in G_k\},\
$$

with $\cup_{k=1}^K G_k = \{1, \ldots, N\}$, $G_k \cap G_j = \emptyset$ and $\sup_{v \in [0,1]} (\|\alpha_k(v) - \alpha_j(v)\|) \neq 0$ for any $k \neq j$, $k = 1, \ldots, K$. The total number of groups K is determined by n_groups.

The predictors are simulated as:

$$
x_{it,j} = 0.2\gamma_i + e_{it,j}, \quad \gamma_i, e_{it,j} \sim i.i.d.N(0,1), \quad j = \{1,\ldots,p\},\
$$

where $e_{it,j}$ denotes a series of innovations. γ_i and e_i are independent of each other.

The errors u_{it} feature a iid standard normal distribution.

In case locations = NULL, the location parameters are drawn from $\sim U[0.3, 0.9]$. In case scales = NULL, the scale parameters are drawn from $\sim U[0.01, 0.09]$. In case polynomial_coef = NULL, the polynomial coefficients are drawn from $\sim U[-20,20]$ and normalized so that all coefficients of one polynomial sum up to 1. The final coefficient function follows as $\alpha_k(t/T) = 3*F(t/T, location, scale) +$ $\sum_{j=1}^d a_j (t/T)^j$, where $F(\cdot, location, scale)$ denotes a cumulative logistic distribution function and a_j reflects a polynomial coefficient.

Value

A list holding

Author(s)

Paul Haimerl

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Examples

```
# Simulate a time-varying panel subject to a time trend and a group structure
sim < sim_t v_DGP(N = 20, n_periods = 50, intercept = TRUE, p = 1)
y \le -\sin\frac{1}{3}y
```
tv_pagfl *Time-varying Pairwise Adaptive Group Fused Lasso*

Description

Estimate a time-varying panel data model with a latent group structure using the pairwise adaptive group fused lasso (*time-varying PAGFL*). The *time-varying PAGFL* jointly identifies the latent group structure and group-specific time-varying functional coefficients. The time-varying coefficients are modeled as polynomial B-splines. The function supports both static and dynamic panel data models.

```
tv_pagfl(
  formula,
  data,
  index = NULL,
  n_periods = NULL,
  lambda,
  d = 3,
  M = \text{floor}(\text{length}(y)^(1/7) - \text{log}(p)),min\_group\_frac = 0.05,
  const_coef = NULL,
  kappa = 2,
  max\_iter = 50000,
  tol_{convergence} = 1e-10,
  tol\_group = 0.001,rho = 0.04 \times \log(N \times n_{\text{periods}})/sqrt(N \times n_{\text{periods}}),
  varrho = 1,
  verbose = TRUE,
  parallel = TRUE,
  ...
)
## S3 method for class 'tvpagfl'
summary(object, ...)
## S3 method for class 'tvpagfl'
formula(x, \ldots)## S3 method for class 'tvpagfl'
df.residual(object, ...)
```

```
## S3 method for class 'tvpagfl'
print(x, \ldots)## S3 method for class 'tvpagfl'
coef(object, ...)
## S3 method for class 'tvpagfl'
residuals(object, ...)
## S3 method for class 'tvpagfl'
fitted(object, ...)
```


Details

Consider the grouped time-varying panel data model

$$
y_{it} = \gamma_i + \beta_i'(t/T)x_{it} + \epsilon_{it}, \quad i = 1, \dots, N, t = 1, \dots, T,
$$

where y_{it} is the scalar dependent variable, γ_i is an individual fixed effect, x_{it} is a $p \times 1$ vector of explanatory variables, and ϵ_{it} is a zero mean error. The coefficient vector $\beta_i(t/T)$ is subject to the latent group pattern

$$
\beta_i\left(\frac{t}{T}\right) = \sum_{k=1}^K \alpha_k\left(\frac{t}{T}\right) \mathbf{1}\{i \in G_k\},\
$$

with $\bigcup_{k=1}^{K} G_k = \{1, \ldots, N\}, G_k \cap G_j = \emptyset$ and $\|\alpha_k - \alpha_j\| \neq 0$ for any $k \neq j, k = 1, \ldots, K$.

The time-varying coefficient functions are estimated as polynomial B-splines using the penalized sieve-technique. To this end, let $B(v)$ denote a $M + d + 1$ vector basis functions, where d denotes the polynomial degree and M the number of interior knots. Then, $\beta_i(t/T)$ and $\alpha_k(t/T)$ are approximated by forming linear combinations of the basis functions $\beta_i(t/T) \approx \pi_i^t B(t/T)$ and $\alpha_i(t/T) \approx \xi'_k B(t/T)$, where π_i and ξ_i are $(M + d + 1) \times p$ coefficient matrices.

The explanatory variables are projected onto the spline basis system, which results in the $(M + d +$ 1) $p \times 1$ vector $z_{it} = x_{it} \otimes B(v)$. Subsequently, the DGP can be reformulated as

$$
y_{it} = \gamma_i + z_{it}' \text{vec}(\pi_i) + u_{it},
$$

where $u_{it} = \epsilon_{it} + \eta_{it}$ and η_{it} reflects a sieve approximation error. We refer to Su et al. (2019, sec. 2) for more details on the sieve technique.

Inspired by Su et al. (2019) and Mehrabani (2023), the time-varying PAGFL jointly estimates the functional coefficients and the group structure by minimizing the criterion

$$
Q_{NT}(\pi, \lambda) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} (\tilde{y}_{it} - \tilde{z}_{it}' \text{vec}(\pi_i))^2 + \frac{\lambda}{N} \sum_{i=1}^{N-1} \sum_{j>i}^{N} \dot{\omega}_{ij} || \pi_i - \pi_j ||
$$

with respect to $\boldsymbol{\pi} = (vec(\pi_i)', \dots, vec(\pi_N)')'.$ $\tilde{a}_{it} = a_{it} - T^{-1} \sum_{t=1}^T a_{it}, a = \{y, z\}$ to concentrate out the individual fixed effects γ_i . λ is the penalty tuning parameter and \dot{w}_{ij} denotes adaptive penalty weights which are obtained by a preliminary non-penalized estimation. ∥ · ∥ represents the Frobenius norm. The solution criterion function is minimized via the iterative alternating direction method of multipliers (*ADMM*) algorithm proposed by Mehrabani (2023, sec. 5.1).

Two individuals are assigned to the same group if $\|\text{vec}(\hat{\pi}_i - \hat{\pi}_j)\| \leq \epsilon_{\text{tol}}$, where ϵ_{tol} is determined by tol_group. Subsequently, the number of groups follows as the number of distinct elements in $\hat{\pi}$. Given an estimated group structure, it is straightforward to obtain post-Lasso estimates ξ using group-wise least squares (see [grouped_tv_plm](#page-5-1)).

We recommend identifying a suitable λ parameter by passing a logarithmically spaced grid of candidate values with a lower limit close to 0 and an upper limit that leads to a fully homogeneous panel. A BIC-type information criterion then selects the best fitting λ value.

In case of an unbalanced panel data set, the earliest and latest available observations per group define the start and end-points of the interval on which the group-specific time-varying coefficients are defined.

Value

An object of class tvpagfl holding

An object of class tvpagfl has print, summary, fitted, residuals, formula, df.residual and coef S3 methods.

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Author(s)

Paul Haimerl

summary(estim)

References

Mehrabani, A. (2023). Estimation and identification of latent group structures in panel data. *Journal of Econometrics*, 235(2), 1464-1482. [doi:10.1016/j.jeconom.2022.12.002.](https://doi.org/10.1016/j.jeconom.2022.12.002)

Su, L., Wang, X., & Jin, S. (2019). Sieve estimation of time-varying panel data models with latent structures. *Journal of Business & Economic Statistics*, 37(2), 334-349. [doi:10.1080/07350015.2017.1340299.](https://doi.org/10.1080/07350015.2017.1340299)

Examples

```
# Simulate a time-varying panel with a trend and a group pattern
set.seed(1)
sim < sim_{rv}DGP(N = 10, n_{peri} = 50, intercept = TRUE, p = 1)df \leftarrow data.frame(y = c(sim\{y}))
# Run the time-varying PAGFL
estim \le tv_pagfl(y \sim ., data = df, n_periods = 50, lambda = 10, parallel = FALSE)
```
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